

COUPON COLLECTION

There are  $n$  coupons. Our goal is to collect all. If  $L$  is the # of cereal boxes I buy to obtain all  $n$  distinct coupons, what's  $E(L)$ ?

The first is always unique, so start at second.  $\frac{(n-1)/n}{1}$

$L_1 = 1, L_2 \sim \text{Geom}(\frac{n-1}{n}), L_j \sim \text{Geom}(\frac{n-j+1}{n})$

$L = L_1 + L_2 + \dots + L_n \rightarrow E(L) = \sum_{j=1}^n E(L_j) = \sum_{j=1}^n \frac{1}{p_j} = \sum_{j=1}^n \frac{n}{n-j+1}$   $k=j-1$

$E(L) = \sum_{k=0}^{n-1} \frac{n}{n-k} = n \sum_{k=0}^{n-1} \frac{1}{n-k} = n \left( \frac{1}{n} + \frac{1}{n-1} + \dots + \frac{1}{2} + 1 \right) = n \ln n + n \gamma + \frac{1}{2}$

$\leftarrow$  gamma harmonic sum  $H_n = \sum_{k=1}^n \frac{1}{k} \approx \ln n + \gamma + \frac{1}{2n}$

$\gamma = 0.5772$ , and this is called Euler's constant.

If  $n=5, E(L) = 5 \ln 5 + 5\gamma + \frac{1}{2} = 11.42 \rightarrow 12$  boxes

VARIANCE

$\sigma_x^2 = \text{Var}(X) = E[(X - E(X))^2] \geq 0 \rightarrow$  standard deviation  $\sigma_x = \sqrt{\text{Var}(X)}$

The units of  $\sigma$  are the same as  $X$ , and  $\sigma^2$  are  $X^2$ .

$\sigma_x^2 = E[X^2 - 2XE(X) + E^2(X)] = \underbrace{E(X^2)}_{\text{2nd moment of X}} + \underbrace{E(-2XE(X))}_{E^2(X)} + \underbrace{E(E^2(X))}_{E^2(X)} = E(X^2) - 2E^2(X) + E^2(X) = E(X^2) - E^2(X)$

If  $X$  is a zero-mean RV (i.e.  $E(X)=0$ ),  $\sigma_x^2 = E(X^2)$ . When is  $\sigma_x^2 = 0$ ? iff  $X = E(X)$

EX What is  $\sigma_x^2$  for a Bernoulli RV?  $\underbrace{E(K^2)}_{\sum_{k=0}^1 k^2 P_K(k)} - \underbrace{E^2(K)}_{p^2} = p - p^2 = p(1-p)$

What kind of a coin would give you maximum (un)certainty?

For certainty, we'd want  $p=1$  so  $\sigma^2=0$ . For uncertainty,  $p=0.5$ .

POISSON RV

Say you have a book with  $n \gg 1$  words, and the probability of each word being a typo is  $0 < p < 1$ . Let  $L$  be the # of typos in the book.

Generally speaking, a large # of Bernoulli trials with low prob. of success.

$P_L(l) = \binom{n}{l} p^l (1-p)^{n-l} \forall l \in \{0, \dots, n\}$   $E(L) = \lambda \rightarrow$  "average arrival rate" =  $np$

CLAIM  
PROOF

As  $n \rightarrow \infty, P_L(l) = \frac{\lambda^l}{l!} \cdot e^{-\lambda}$

$P_L(l) = \frac{n!}{l!(n-l)!} \left(\frac{\lambda}{n}\right)^l \left(1 - \frac{\lambda}{n}\right)^{n-l} \left(1 - \frac{\lambda}{n}\right)^l = \frac{n(n-1)\dots(n-l+1)}{l! n^l} \lambda^l \left(1 - \frac{\lambda}{n}\right)^{n-l} p = \frac{\lambda^l}{l!} \left(1 - \frac{\lambda}{n}\right)^{n-l}$

$= \frac{\lambda^l}{l!} \cdot \underbrace{\frac{n \cdot (n-1) \cdot \dots \cdot (n-l+1)}{n^l}}_{1 \cdot 1 \cdot \dots \cdot 1} \left(1 - \frac{\lambda}{n}\right)^{n-l} = \frac{\lambda^l}{l!} \cdot e^{-\lambda}$

from calculus,  $\lim_{x \rightarrow \infty} (1+x)^{1/x} = \lim_{y \rightarrow \infty} (1+1/y)^y = e$



## MEAN OF POISSON RV

$$E(L) = \sum_{l=0}^{\infty} l \frac{\lambda^l}{l!} e^{-\lambda} = \lambda$$

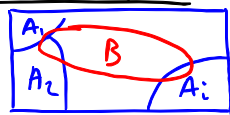
we want  $\lambda = np$

$$= \sum_{l=1}^{\infty} \frac{\lambda^l}{(l-1)!} e^{-\lambda} = \lambda \left[ \sum_{l=1}^{\infty} \frac{\lambda^{l-1}}{(l-1)!} \right] e^{-\lambda} = \lambda \cdot e^{\lambda} \cdot e^{-\lambda} = \lambda$$

## VARIANCE OF POISSON RV

Also  $\lambda$ .

## JOINT PMFs



$$P(B) = \sum_{i=1}^{\infty} P(B \cap A_i)$$

Marginal PMF  $P_X(x) = P(X=x)$

$$P_{X,Y}(x,y) = P_r(X=x \cap Y=y)$$

$$P_{X|Y}(x|y) = \frac{P_r(X=x \cap Y=y)}{P_Y(y)} = \frac{P_{X,Y}(x,y)}{P_Y(y)}$$

Plotting this in 2D is hard. Anything higher is hopeless.